

Auteurs

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① FRAMEWORK

Goal. Sample « according » to a *target distribution* π on \mathbb{X} that does not work well with MCMC algorithms.

Idea. Transform a sample $(\tilde{X}_k)_{k=1..K}$ from an *instrumental distribution* $\tilde{\pi}$ into a sample $(X_i)_{i=1..I}$ distributed under π .

Known techniques and drawbacks.

- *Rejection sampling*: need to know a constant M such that $\pi \leq M\tilde{\pi}$.
- *Importance sampling*: doesn't give a « real » sample but just a weighted sample used to compute estimates.

Markov chain approach. Instrumental and target samples do not have to be i.i.d. but can be generated by Markov chains targeting π and $\tilde{\pi}$ resp.

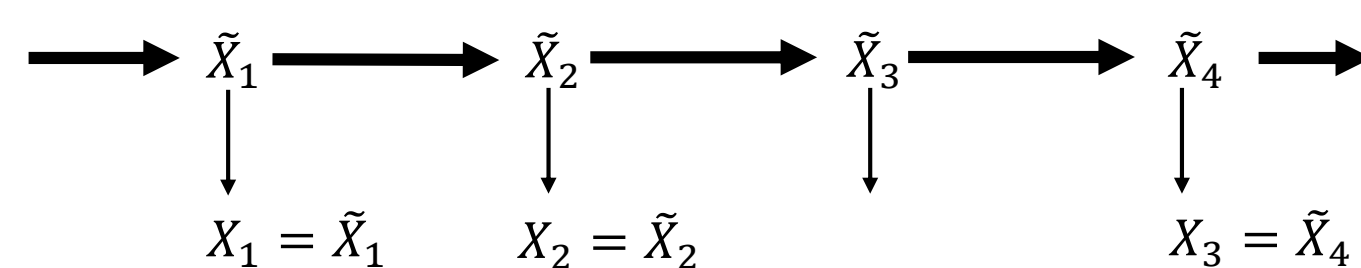
Notations. Q : Markov kernel targeting $\tilde{\pi}$, $\rho_\kappa = \kappa \frac{d\pi}{d\tilde{\pi}}$ density ratio, κ : tuning param.

② PRESENTATION OF THE IMC

REJECTION MC

Constraint. Constant M such that $\pi \leq M\tilde{\pi}$... Problem if such a const. M does not exist, is unknown or is too large leading to excessive rejection rate.

Algo. Perform the rejection algorithm on a sample $(\tilde{X}_k)_{k \in \mathbb{N}}$ generated by Q with acceptance function $\rho = \rho_M$.

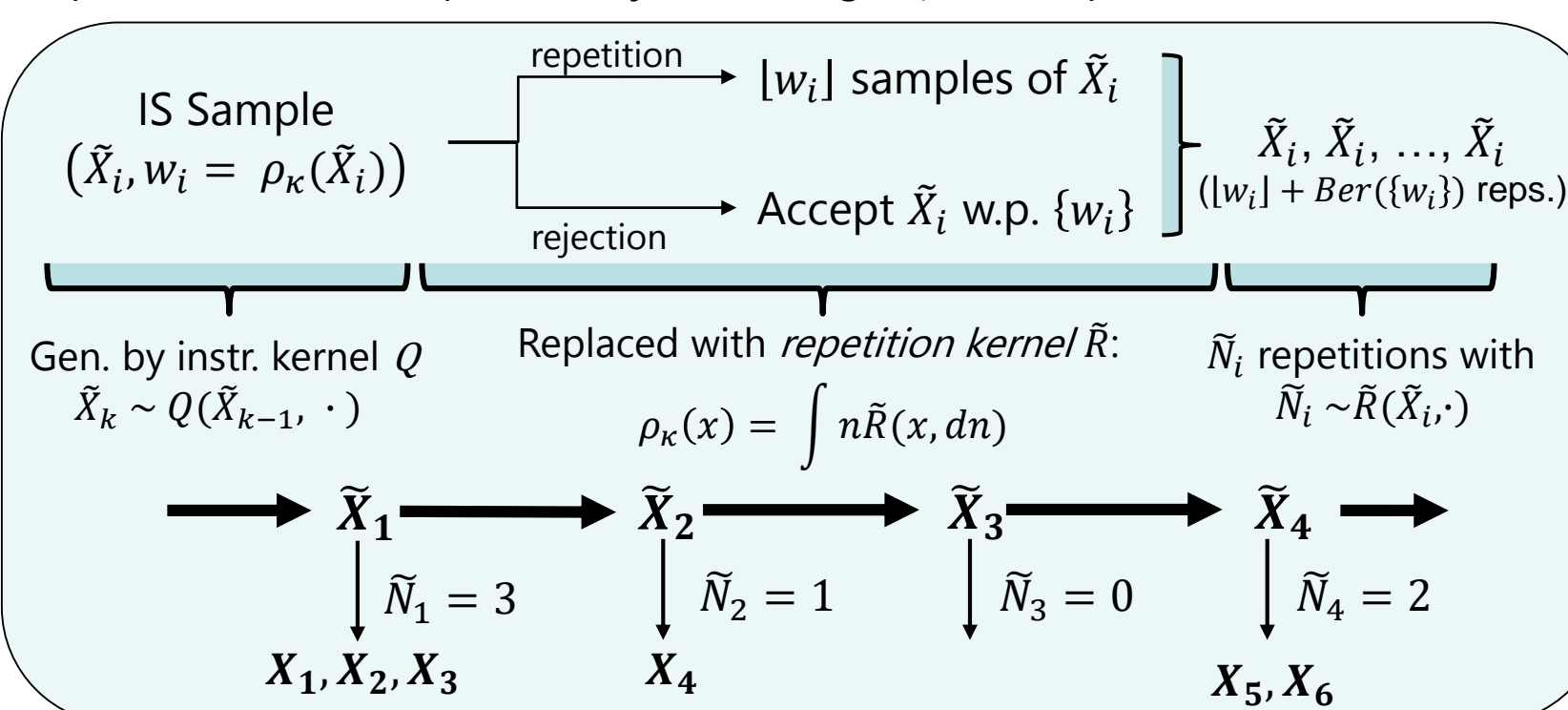


Output. A *rejection chain* $(X_i)_{i=1..I}$, generated by the submarkovian *rejection kernel* S defined, for $h \in F_+(\mathbb{X})$, by:

$$Sh(x) = \sum_{k=1}^{\infty} \mathbb{E}_x^Q \left[h(\tilde{X}_k) \rho(\tilde{X}_k) \prod_{i=1}^{k-1} (1 - \rho(\tilde{X}_i)) \right]$$

HEURISTICAL IMC

Idea. Remove constraint $\pi \leq M\tilde{\pi}$ i.e. allowing $\rho_\kappa \geq 1$ using repetitions in conjunction with the Rejection MC algorithm. Repetitions allow us to perform rejection using $\{\rho_\kappa\}$ as acceptance function.



Choice of \tilde{R} . The optimal choice for \tilde{R} in terms of variance reduction is the *shifted Bernouilli*:

$$\tilde{R}_{opt}(x, \cdot) = (1 - \{\rho_\kappa(x)\})\delta_{\{\rho_\kappa(x)\}} + \{\rho_\kappa(x)\}\delta_{\{\rho_\kappa(x)\}+1}$$

A **MH algorithm** with proposal $A(x, dy)$ and acceptance rate $\alpha(x, y)$ targeting π can be obtained from the IMC by taking $\tilde{R}(x, \cdot) = \text{Geom}(p(x))$ where $p(x) = \int_{\mathbb{X}} \alpha(x, y) A(x, dy)$.

④ NUMERICAL EXPERIMENTS

IMPROVED SAMPLER

Question. Can IMC improve the effectiveness of a sampler targeting a multimodal distribution by using it on an instrumental smoothed version of the latter distrib. ?

Setup. We target $\pi = \sum_{i=1}^n \mathcal{N}(\mu_i, I_d)$ an unnormalized gauss. mix., with $\mu_i \sim \mathcal{N}(0, 10^2 I_d)$

- $d = 5, n = 6$
 - $\tilde{\pi} = \pi^\beta$ for $\beta \in (0, 1)$ and $(x, \cdot) = \tilde{\pi}(\cdot)$ for all $x \in \mathbb{X}$
 - Q is a No-U-turn Sampler (NUTS) and \tilde{R} is a shifted Bernouilli kernel
- We estimate the MSE of π by running 200 chains for each $\beta \in \{0.004, 0.01, 0.04, 0.1, 1\}$

β	0.004	0.010	0.040	0.100	1.000
MSE	16.150	6.123	0.544	17.863	33.982

INDEP. IMC VS. INDEP. MH ALGORITHM

Setup. We target $\pi = \frac{1}{2}\mathcal{N}(0, 4) + \frac{1}{2}\mathcal{N}(0, 4)$ using :

- $\tilde{\pi} = \mathcal{N}(0, 4)$
- $Q(x, \cdot) = \tilde{\pi}(\cdot)$ for all $x \in \mathbb{X}$
- \tilde{R} a shifted Bernouilli kernel
- $N_{iterations} = 100,000$

③ THE IMC ALGORITHM

EXTENDED MARKOV CHAIN

Setup. Add remaining repetition counter as second component and extract the sample from the first component at the end.

Property. Extended chain $(X_i, N_i)_{i \in \mathbb{N}}$ is markovian with transition kernel :

$$Ph(x, n) = \mathbb{1}_{\{n \geq 1\}} h(x, n-1) + \mathbb{1}_{\{n=0\}} \sum_{n' \geq 1} \int_{\mathbb{X}} S(x, dx') \frac{\tilde{R}(x', n'+1)}{\rho_{\tilde{R}}(x')} h(x', n')$$

where $\rho_{\tilde{R}}(x') = \tilde{R}(x, [1: \infty))$.

Interpretation. While the counter is nonzero, decrease it by one. When it reaches zero :

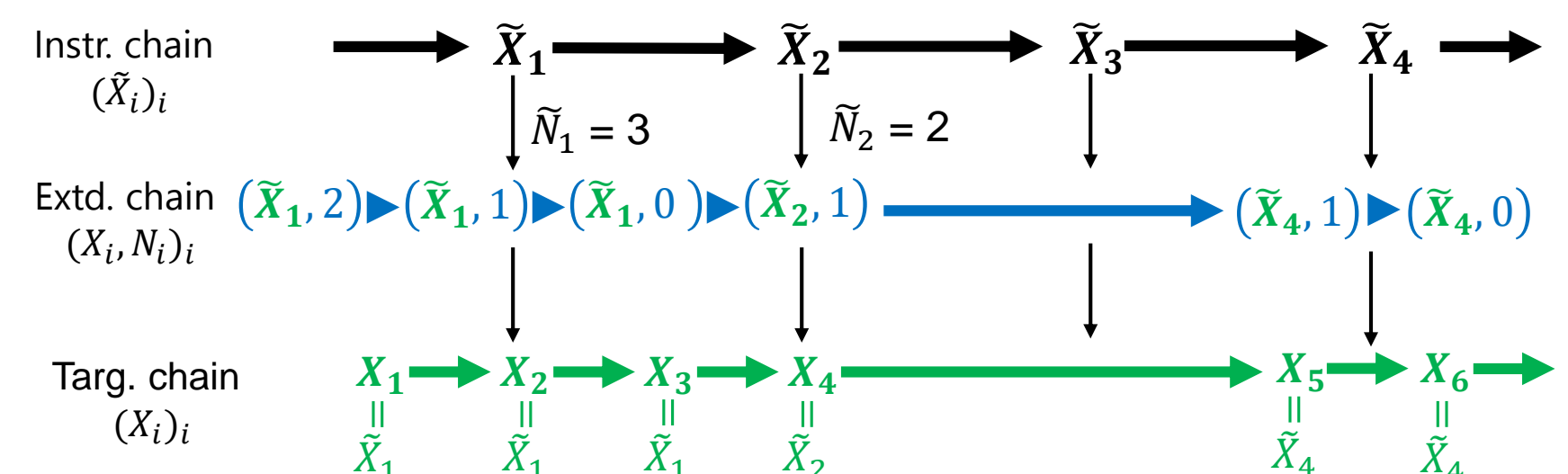
- Use transition kernel S ensuring transitions from one *accepted* point to another
- Use *conditionned* transition kernel $R(x, n) = \frac{\tilde{R}(x, n+1)}{\rho_{\tilde{R}}(x)}$ taking into account that we only draw accepted points i.e. $\tilde{N} \geq 1$.

ALGORITHM AND MAIN RESULTS

INIT. Set an arbitrary \tilde{X}_0 and $i = 0$.

REPEAT.

1. Draw $\tilde{X}_k \sim Q(\tilde{X}_{k-1}, \cdot)$ and $\tilde{N}_k \sim \tilde{R}(\tilde{X}_k, \cdot)$
2. Set $N_i = \tilde{N}_k$
3. While $N_i \geq 1$: $(X_i, N_i) \leftarrow (\tilde{X}_k, N_i - 1)$; $i \leftarrow i + 1$;



Proposition [Invariant measure]. P admits an invariant probability measure $\bar{\pi}$, which has π as its first marginal. Moreover, $\bar{\pi}$ is unique (under mild additional conditions).

Theorem [SLLN]. If Q admits a SLLN starting from any initial distribution, i.e. for any $\xi \in M_1(\mathbb{X})$ and g measurable function such that $\bar{\pi}(g) < \infty$,

$$\lim_{n \rightarrow \infty} n^{-1} \sum_{k=0}^{n-1} g(\tilde{X}_k) = \bar{\pi}(g), \quad \mathbb{P}_\xi^Q - a.s.$$

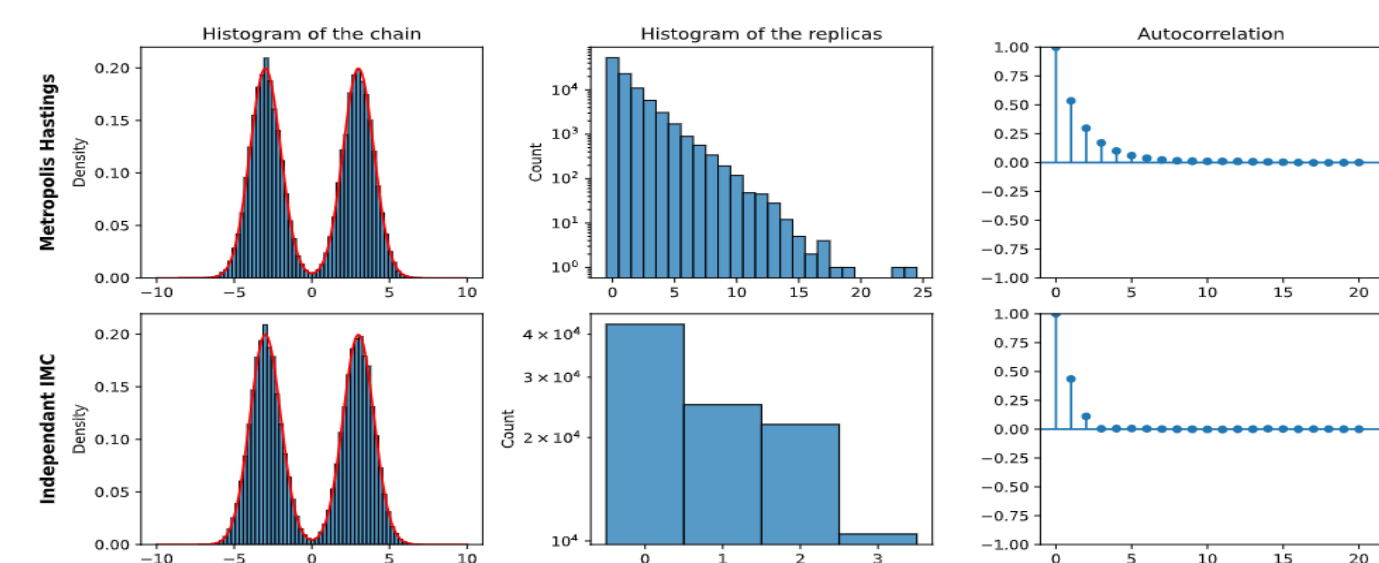
Then P also admits a SLLN starting from any initial distribution on $\mathbb{X} \times \mathbb{N}$.

Theorem [CLT]. Assume that Q admits a solution to the Poisson equation associated to $\rho_\kappa(h - \pi(h))$. Then, under mild additional assumptions on \tilde{R} , the kernel P admits a CLT for h , i.e. there exists a constant $\sigma^2(h) > 0$ and a distribution χ , easily expressed using ξ , and corresponding to the distribution of the first accepted couple, such that,

$$n^{-1/2} \sum_{i=1}^n (h(X_i) - \pi(h)) \rightsquigarrow^\mathcal{L} \mathcal{N}(0, \sigma^2(h)), \quad \mathbb{P}_\xi^P - \text{law.}$$

Theorem [Geometric ergodicity]. Assume a specific set is a small set for Q and a drift condition on that set for a function V . Then there exist constants $\delta, \beta_r > 1$ and $\zeta < \infty$ such that for all $\xi' \in M_1(\mathbb{X} \times \mathbb{N})$,

$$\sum_{k=1}^{\infty} \delta^k d_{TV}(\xi P^k, \bar{\pi}) \leq \zeta \int_{\mathbb{X} \times \mathbb{N}} \beta_r^n V(x) \xi'(dx dn)$$



Analysis. Both histograms of the chain look similar, but autocorrelation suggests better mixing for IMC. Also fewer max repetitions for IMC (reminder : geometric distrib. for MH).

	X	X^2	X^3	X^4
MH	6.20e-03	2.33e-02	1.49e+00	1.57e+01
IMC	3.49e-03	9.74e-03	8.40e-01	7.18e+00

Table. MSE for the first 4 moments for 10^4 repetitions of chains of length 10^4 .